R20-Exploratory Factor analysis and principal component analysis in R Colleen F. Moore Feb 2015 cfmoore@wisc.edu Prof Emerita, University of Wisconsin-Madison Affiliate Professor, Montana State University, Bozeman

In R there are several ways to do exploratory factor and principal components analysis. Best reference, and developer of the 'psych' package: William Revelle, see links inside R in documentation on the 'psych' package. Ch 6 of his forthcoming book is highly recommended.

Also very good, Michael Friendly's page, not specific to R: http://www.psych.yorku.ca/lab/psy6140/fa/facplan.htm

My handout here is not intended to be a lecture handout, but a relatively quick reference for 'how to' in R.

Contents of this handout:

I. Preliminaries (test correlation matrix, find SMC, look for outliers)II. Principal components analysis (two options, princomp or principal). Scree plots.III. Factor analysis ('factanal' or 'fa')

- IV. Other nifty things in the 'psych' package, including Very Simple Structure, parallel analysis (both help choose number of factors to fit), comparing factor analyses across samples or within sample, Kaiser-Meyer-Olin index of sampling adequacy, Cronbach's alpha
- V. Other nifty thing (from me). How to randomly split a sample in two to test sample separately.

> library(psych) ## bring the psych package into R memory, for a lot of what is done below

I. Preliminaries (and how to do them) before diving into principal components or factor analysis

A. Test to see if your correlation matrix differs significantly from the identity matrix. You don't want to be fitting just error. See section IV.A.1. below.
B. Do you have a reasonable set of measures, or do some items not belong in this analysis? Find the squared multiple correlations (smc) of each variable with the others. Inspect for low values, read the items that have low smc values, and decide whether to remove them. See section IV.A.2 below. If you are constructing a new scale, you will want to remove items after fitting a model also.
C. Look for outliers using Mahalanobis distances (D2):

> outlier(asiq, plot=T, bad=10, na.rm=T) # in psych package

In a large sample, ask it to flag more bad values than in a smaller sample. Also, Mahalanobis distances are supposed to be distributed as a chi-squared distribution, with df = number of variables going into the distance calculation. Can get some idea about how far out of your distribution the outliers are by looking at the p-values of the chisq distribution. For my 8 variable example below as follows: > qchisq(.01, 8, lower.tail=F) # p=.01, df=8, we want the upper tail [1] 20.09024 ## this says that if an outlier has a distance over 20, it is in the upper 1% of distribution. > pchisq(15, 8, lower.tail=F) ## this gives you the prob of a given chisq value [1] 0.05914546

II. Principal components
There are at least two ways to calculate principal components in R.
A. princomp - does principal components, yields eigenvalues.
 Minimal output, can't control # of components??

> pca2b=princomp(mat2,factors=2) ## data are in mat2 > summary(pca2b) Importance of components:

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Comp.1 Comp.2 Comp.3 Comp.4 Comp.5 Comp.6 Standard deviation 5.4619755 3.5208939 1.47284209 1.21102164 1.03527649 1.00654155 Proportion of Variance 0.6106818 0.2537589 0.04440459 0.03002059 0.02193957 0.02073857 Cumulative Proportion 0.6106818 0.8644407 0.90884532 0.93886592 0.96080549 0.98154406 Comp.7 Comp.8 0.84603547 0.431089184 Standard deviation Proportion of Variance 0.01465186 0.003804081 Cumulative Proportion 0.99619592 1.00000000

look at scree plot, there are 2 ways to do this, note difference in scaling, the
output immediately below is the square of the sd of the components shown above, the
plot on the next page has the sd's themselves.
> screeplot(pca2b);

· Selection (peals),



From inside the 'psych' package can also make scree plot. Notice different scaling
from above. One is the square root of the other.
> VSS.scree(mat2)



B. Can calculate principal components using 'principal' in 'psych' package > pca2=principal(mat2, nfactors=2, rotate="varimax", scores=F) > pca2 ## get the output from R Principal Components Analysis Call: principal(r = mat2, nfactors = 2, rotate = "varimax", scores = F) Standardized loadings (pattern matrix) based upon correlation matrix RC1 RC2 h2 u2 Researchers announce 0.16 0.90 0.84 0.157 Researchers communicate quickly 0.05 0.93 0.87 0.134 Researchers_pos_contribute 0.39 0.77 0.75 0.251 Researchers available 0.25 0.88 0.84 0.156 Influence worthwhile 0.92 0.23 0.90 0.097 Influence_benefits_community 0.92 0.18 0.89 0.113 Influence_important_topic 0.96 0.22 0.96 0.039 Influence healthcare 0.93 0.14 0.89 0.110 RC1 RC2

 SS loadings
 3.74
 3.21

 Proportion Var
 0.47
 0.40

 Cumulative Var
 0.47
 0.87

 Proportion Explained
 0.54
 0.46

 Cumulative Proportion
 0.54
 1.00

Test of the hypothesis that 2 components are sufficient.

The degrees of freedom for the null model are 28 and the objective function was 9.3 The degrees of freedom for the model are 13 and the objective function was 0.7 The total number of observations was 70 with MLE Chi Square = 44.89 with prob < 2.2e-05

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Fit based upon off diagonal values = 1 ## get more output > pca2\$loadings Loadings: RC1 RC2 Researchers_announce 0.157 0.905 Researchers_communicate_quickly 0.929 Researchers_pos_contribute 0.392 0.772 Researchers_available 0.254 0.883 Researchers_available0.2540.883Influence_worthwhile0.9230.225Influence_benefits_community0.9240.183Influence_important_topic0.9560.216Influence_healthcare0.9330.137 RC1 RC2 SS loadings 3.736 3.207 Proportion Var 0.467 0.401 Cumulative Var 0.467 0.868 ## See documentation for how to get residuals, scores, and other rotations. Notice that asking for the loadings stored, the 'principal' program in 'psych' package omits loadings below a low cutoff value. III. "Common factors" or true factor analysis A. Can use fa in psych package > paf2=fa(mat2,nfactors=2,rotate="varimax",SMC=T,symmetric=T, fm="pa") > paf2 ## ask for results Factor Analysis using method = pa Call: fa(r = mat2, nfactors = 2, rotate = "varimax", SMC = T, symmetric = T, fm = "pa") Standardized loadings (pattern matrix) based upon correlation matrix Researchers_announce PA1 PA2 h2 u2 com 0.16 0.87 0.79 0.212 1.1 Researchers_communicate_quickly 0.06 0.90 0.82 0.178 1.0

 Researchers_communicate_quickiy
 0.06 0.90 0.62 0.176 1.0

 Researchers_pos_contribute
 0.38 0.72 0.67 0.334 1.5

 Researchers_available
 0.26 0.86 0.80 0.203 1.2

 Influence_worthwhile
 0.90 0.23 0.87 0.131 1.1

 Influence_benefits_community
 0.90 0.19 0.84 0.160 1.1

 Influence_important_topic
 0.97 0.22 0.98 0.017 1.1

 Influence_healthcare
 0.90 0.15 0.84 0.160 1.1

 PA1 PA2
 SS loadings
 3.62
 2.99

 Proportion Var
 0.45
 0.37

 Cumulative Var
 0.45
 0.83
 Proportion Explained 0.55 0.45 Cumulative Proportion 0.55 1.00 Mean item complexity = 1.1Test of the hypothesis that 2 factors are sufficient. The degrees of freedom for the null model are 28 and the objective function was 9.3 with Chi Square of 609.43 The degrees of freedom for the model are 13 $\,$ and the objective function was $\,$ 0.44 $\,$ The root mean square of the residuals (RMSR) is 0.02 The df corrected root mean square of the residuals is 0.02

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The harmonic number of observations is 70 with the empirical chi square 1.13 with prob < 1 The total number of observations was 70 with MLE Chi Square = 28.15 with prob < 0.0086 Tucker Lewis Index of factoring reliability = 0.943 RMSEA index = 0.139 and the 90 % confidence intervals are 0.062 0.195BIC = -27.08Fit based upon off diagonal values = 1 Measures of factor score adequacy PA1 PA2 Correlation of scores with factors 0.99 0.96 Multiple R square of scores with factors 0.99 0.92 Minimum correlation of possible factor scores 0.98 0.85 ## See documentation for other options for both rotation and factoring methods. B. Another option: factanal, which does maximum likelihood factor analysis > mlf2=factanal(mat2, factors=2, rotation="varimax"); > mlf2; ## get R to show results Call: factanal(x = mat2, factors = 2, rotation = "varimax") Uniquenesses: Researchers announce Researchers communicate quickly 0.198 0.172 Researchers_available Researchers_pos_contribute 0.348 0.221 Influence_worthwhile Influence_benefits_community 0.169 0.147 Influence important topic Influence healthcare 0.010 0.138 Loadings: Factor1 Factor2 Researchers_announce 0.147 0.883 Researchers_communicate_quickly 0.908 Researchers_pos_contributtoResearchers_available0.2480.847Influence_worthwhile0.8780.245Influence_benefits_community0.9020.197Influence_important_topic0.9680.228Influence_bealthcare0.9150.155 Researchers_pos_contribute 0.364 0.721 Researchers_available 0.248 0.847 Factor1 Factor2 SS loadings 3.580 3.016 Proportion Var 0.447 0.377 Cumulative Var 0.447 0.824 Test of the hypothesis that 2 factors are sufficient. The chi square statistic is 25.15 on 13 degrees of freedom. The p-value is 0.0221 ## See documentation for estimating factor scores, etc

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IV. Other nifty things related to principal components or factor analysis in psych package A. Bartlett's test for a correlation matrix (is it identity matrix + error). You shouldn't do factor analysis on a random matrix. Also known as Bartlett's test of sphericity. You want the Bartlett test to have a small p-value. > cortest.bartlett(D2AxS[,26:31]) ## example using columns 26:31 of my data R was not square, finding R from data \$chisq [1] 75.47375 \$p.value [1] 4.648804e-10 Śdf [1] 15 > cortest.mat(D2AxS[,26:31]); ## also calculates Bartlett's test Bartlett's test of is R = ITests of correlation matrices Call:cortest.mat(R1 = D2AxS[, 26:31]) Chi Square value 75.47 with df = 15 with probability < 4.6e-10Warning message: In cortest.mat(D2AxS[, 26:31]) : R1 matrix was not square, correlations found> cortest.jennrich compares matrices > cortest.normal; ## differs but can use this to compare pairs of matrices, which is interesting to do if you have two samples tested on the same variables B. Get the squared multiple correlations of each variable with all the others.

Look at these to see if you should throw out some variables. Some say use a .30 (about 10% shared variance) criterion, but it is just a rule of thumb. > smc1= smc(D2AxS[,26:31]); ## a few columns of my data again subnig vta acmb amyg bs caud 0.7887606 0.6336429 0.9430341 0.8579995 0.6595345 0.9502663

When you use this, it is important to look at the content of the items and to think. You can also plot the cumulative distribution function of the squared multiple correlations and look at it to get a feel for whether some items don't correlate very well with the others.

> plot.ecdf(smc1, main="Some Brain Areas, Sq mult corrs", xlab="sq mult corr")



> abline(v=.3, lty=2) ## puts a vertical line at .30 to aid eyeball, which wasn't necessary in this case. Some recommend removing items that have smc less than .3 with other items. Can position it wherever you like.

- C. VSS, or Very Simple structure, an aid to choosing number of factors.
 - VSS gives a lot of information, not just the VSS fit, which is also plotted. Complexity is the number of factors that each variable loads on. Choose the number of factors based on the maximum value of the VSS. Note, this gives a different answer for the data example below from the parallel analysis in the next section. This also included Velicer's MAP function, which should be minimized.

```
> VSS(asiq, n=8, rotate="varimax", fm="pa", plot=T)
# different data set from earlier in this handout
```

```
The Velicer MAP achieves a minimum of 0.01 with 5 factors
BIC achieves a minimum of -3054.04 with 8 factors
Sample Size adjusted BIC achieves a minimum of -474.89 with 8 factors
```

Statistics by number of factors

```
prob sqresid fit RMSEA
  vssl vss2
              map dof chisq
                                                        BIC SABIC complex eChisq
                                     68 0.57 0.099
1 0.57 0.00 0.0207 1127 13383 0.0e+00
                                                       5451 9031
                                                                    1.0 38321
2 0.64 0.74 0.0130 1079 8950 0.0e+00
                                         42 0.74 0.081 1356
                                                            4783
                                                                     1.3
                                                                          16478
                       6313 0.0e+00
3 0.60 0.77 0.0087 1032
                                         30 0.81 0.068
                                                       -950
                                                             2328
                                                                     1.5
                                                                           7918
                       4948 0.0e+00
4 0.58 0.78 0.0073 986
                                         25 0.84 0.060 -1991
                                                            1141
                                                                           4972
                                                                     1.7
5 0.50 0.74 0.0069 941
                       4068 0.0e+00
                                         22 0.86 0.055 -2555
                                                                           3454
                                                             434
                                                                     2.0
6 0.50 0.74 0.0073 897
                       3613 2.6e-321
                                        20 0.87 0.052 -2700
                                                             149
                                                                     2.1
                                                                           2854
7 0.49 0.74 0.0076 854
                       3071 7.2e-247
                                         19 0.88 0.048 -2940
                                                             -227
                                                                           2267
                                                                     2.1
8 0.48 0.74 0.0076 812 2661 5.8e-195
                                             17 0.89 0.045 -3054 -475
    1753
2.1
   SRMR eCRMS eBIC
1 0.120 0.122 30389
2 0.078 0.082 8884
3 0.054 0.058
               654
4 0.043 0.047 -1967
5 0.036 0.040 -3169
6 0.033 0.037 -3459
```

```
7 0.029 0.034 -3743
8 0.026 0.031 -3962
```



D. **Parallel analysis.** Choose the number of factors by simulating a random data set, and choosing the point where the eigenvalues of the real data fall below the simulated data.

> pfa3=fa.parallel(asiq, fm="minres",fa="both")

Parallel analysis suggests that the number of factors = 10 and the number of components = 8

Parallel Analysis Scree Plots



```
E. Compare factor solutions. Some writers say to "Factor the data by several different analytical procedures and hold sacred only those factors that appear across all the procedures used." (Gorsuch, Factor Analysis, p. 330, 1983).
> pcal=principal(asiq, nfactors=5, rotate="promax", scores=F)
> paf1=fa(asiq,nfactors=5,rotate="varimax",SMC=T,symmetric=T, fm="pa")
> factor.congruence(pcal,paf1); # compare princ comp and factor analysis
PA1 PA3 PA2 PA4 PA5
PC3 -0.39 -0.98 0.12 0.02 -0.05
PC1 0.94 0.36 -0.09 0.00 0.24
PC2 -0.08 -0.03 0.97 -0.13 0.07
PC4 0.06 -0.04 -0.23 0.98 0.18
PC5 0.14 -0.04 0.01 -0.06 0.94
```

For this example, models are fit with 5 principal components or with 5 factors, and different rotations are applied. I have highlighted the diagonal elements, because the components/factors are not ordered the same.

Another way to do this is to use the solution from one set of data and apply it to another (for example, a random half of the sample). > predict.psych; (see documentation in R)

```
G. Sort the factors by loading size, makes it easier to think through.
   > fa.sort(faresults) ## where `faresults' has the results of a factor analysis
   > fa.organize (faresults) ## leaves items in original order
H. ..
I. Kaiser-Meyer-Olin test of ``sampling adequacy". Some say don't extract factors
   if this is below .50. The higher the better.
        > KMO(D2AxS[, 26:31]) ## a few columns of a small data set
        Kaiser-Meyer-Olkin factor adequacy
        Call: KMO(r = D2AxS[, 26:31])
        Overall MSA = 0.55
        MSA for each item =
        subnig vta acmb amyg bs caud
        0.58 0.49 0.57 0.58 0.47 0.56
```

```
J. Calculate Cronbach's alpha (see Revelle's documentation for other methods that
   are less entrenched but perhaps better )
   First make a matrix with the items in your scale. Then use 'alpha'.
   > library(psych); ## just a reminder to you to activate the 'psych' package
   > fac1=data.frame(R_understandable_language, R_friendly, R_available,
            R announceresults, R reportresults, R sigcontribution community,
            R sigcontribution personal) # put the variables in a data frame
   > alphafac1=alpha(fac1,keys=NULL, cumulative=F,na.rm=T)
   > alphafac1 ## get the results from R
   Reliability analysis
   Call: alpha(x = fac1, keys = NULL, cumulative = F, na.rm = T)
      raw alpha std.alpha G6(smc) average r S/N ase mean sd
            0.9
                      0.9 \quad 0.92 \quad 0.5\overline{6} \ 8.8 \ 0.034 \ 6.5 \ 1.9
    lower alpha upper 95% confidence boundaries
    0.83 0.9 0.96
    Reliability if an item is dropped:
                                    raw_alpha std.alpha G6(smc) average r S/N alpha se
   R understandable_language
                                      - 0.89 0.89 0.91 0.\overline{58} 8.4 0.038
   R_anderseandable_language0.090.090.090.000R_friendly0.890.890.900.588.20.038R_available0.880.880.910.567.50.040R_announceresults0.870.870.890.536.80.041R_reportresults0.870.870.890.546.90.041R_sigcontribution_community0.870.880.900.547.10.041R_sigcontribution_personal0.890.890.910.588.20.039
    Item statistics
                                      n r r.cor r.drop mean sd
   R understandable language 112 0.72 0.66 0.60 6.7 2.3
   R_friendly 113 0.73 0.68 0.61 7.5 2.1
   R_available114 0.79 0.75 0.70 6.4 2.3R_announceresults112 0.86 0.86 0.80 6.3 2.4R_reportresults115 0.85 0.85 0.79 6.3 2.5
   R_sigcontribution_community 114 0.83 0.81 0.78 6.5 2.7
R_sigcontribution_personal 114 0.74 0.68 0.64 5.8 2.9
```

V. Nifty stuff (Not inside 'psych' package)

Sometimes we want to split a large sample in order to cross validate a factor solution.

```
K. Code to split a large enough data set randomly into 2 groups (won't be exactly
   equal, but fiddle around until the split is close to equal)
   > x=as.matrix(sample(c(0,1),1139, replace=T))
   # the data sample has 1139 observations, so create a variable, x, with 1139
   randomly sampled 1's and 0's.
   > mean(x) # find the mean to see how close to an equal split it was
   [1] 0.4978051 ## can re-do the split until we get one that is about 50-50
   > newdat=cbind(x,asiq); ## column bind the new variable with original data
   > ncol(newdat) ## original data had 49 columns, checking that now I have 50
   [1] 50
   > newdat1=subset(newdat, x==1); ## now I extract the cases with the 1's
   ## the '==' means logically true
   > nrow(newdat1);
   [1] 567 ## there are 567 observations in the data set labeled 1.
   > newdat0=subset(newdat, x==0); ## extract the cases with the 0's
   > nrow(newdat0); ## check the number of observations
```

[1] 572 ## there are 572 observations in the data set labeled 0. > library(MASS) ## the write.matrix function is in MASS package > write.matrix(newdat0,file="TA0data.txt",sep=" ") # save the results for the cases with 0's Instead of using a blank as the separator you can use a comma to create a csv file > write.matrix(newdat1, file="TA1data.txt",sep=" ") # save the results for the cases with 1's.

- L. Code to do what M. more nifty code maybe
- 1. (blah blah to be continued... perhaps) 2.